

# Identification of Nonignorable Nonresponse Mechanisms for Two Way Contingency Tables

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## Abstract

In this paper, we use decomposable graphical models to describe the mechanisms of nonignorable nonresponse in contingency tables classified by two binary variables, and we discuss identification of parameters in these models. For an unidentifiable model, we propose adding covariates which are always observed such that this model becomes identifiable.

## 1. Introduction

Contingency tables with nonresponses have been discussed by many authors. According to the terminology of Little and Rubin (1987), the nonresponse mechanisms can be classified into three types: missing completely at random (MCAR), missing at random (MAR) and nonignorable nonresponse mechanisms. The MCAR and MAR nonresponse mechanisms are also said to be ignorable. Fay (1986) discussed nonignorable nonresponse mechanisms and used an indicator for each variable which is subjected to nonresponse. Baker and Laird (1988) developed a log linear model for categorical variables subject to nonresponse. Baker (1992) provided a class of models of nonresponse mechanisms with closed form estimates of cell probabilities. Molenberghs et al. (1999a, b) discussed some issues on models with incomplete categorical data, including identifying of parameters and so on. Glonek(1999) presented necessary and sufficient conditions of global identifiability for simple nonignorable nonresponse models with one or two binary responses. In this paper, we use decomposable graphical models to describe the mechanisms of nonignorable nonresponse in contingency tables classified by two binary variables, and we discuss identification of parameters in these models. For an unidentifiable model, we propose adding covariates which are always observed such that this model becomes identifiable.

Section 2 introduces some notations. Section 3 discusses identification of nonresponse mechanisms for  $2 \times 2$  contingency tables subject to nonresponses. Section 4 discuss estimates of the parameters.

## 2. Notation

Let  $(Y_1, Y_2)$  denote the response vector with  $Y_t = 0$  or  $1$  for  $t = 1, 2$ . We introduce a response indicator  $R_t$  with value 1 if the response  $Y_t$  is obtained and 0 otherwise. We use decomposable graphical models to describe the association among the four variables:  $Y_1$ ,

Table 1: Observed data subject to nonresponses.

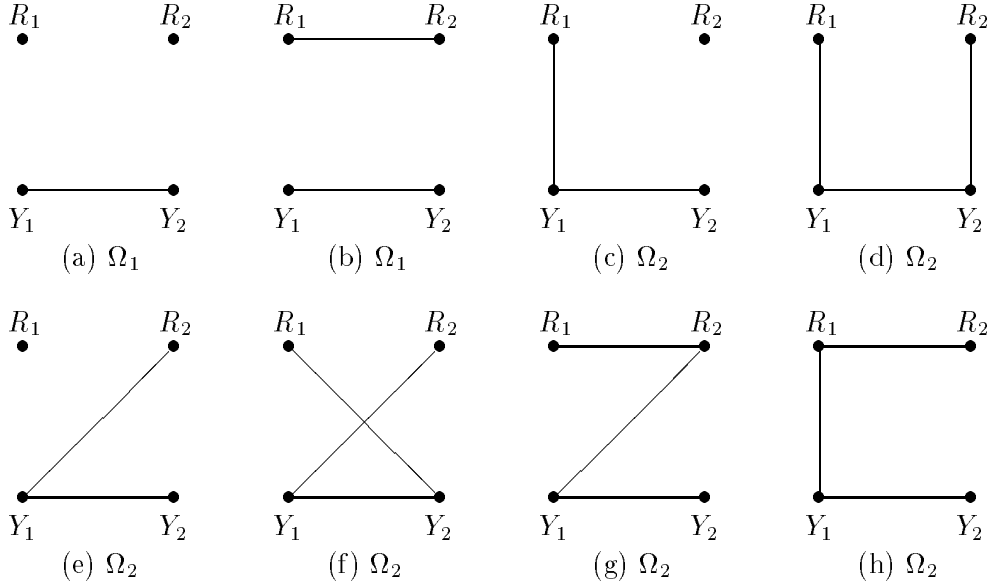
$n_{0011}$	$n_{0111}$	$n_{0+10}$
$n_{1011}$	$n_{1111}$	$n_{1+10}$
$n_{+001}$	$n_{+101}$	$n_{++00}$

$Y_2$ ,  $R_1$  and  $R_2$ . Let  $G = (V, E)$  denote an undirected graph where  $V = \{Y_1, Y_2, R_1, R_2\}$  is the set of nodes and  $E$  are the set of undirected edges between these nodes. The absence of an edge between a pair of nodes means that the corresponding variables in this pair are independent conditionally on the other variables. A graph is decomposable if it does not has any cycle of length than or equal to 4 without a chord. A decomposable graph denotes a decomposable graphical model. Let  $M(G)$  denote the set of all possible joint probabilities based on the graphical model  $G$ .

Observed data can be denoted as Table 1, where  $n_{ijkl}$  denotes the observed frequency for  $Y_1 = i$ ,  $Y_2 = j$ ,  $R_1 = k$  and  $R_2 = l$ , and ‘+’ denotes that the corresponding variable is missing. For example,  $n_{1+10}$  denotes the frequency for  $Y_1$  is observed with value 1 but  $Y_2$  is missing.

### 3. Identifiability of Parameters

If any joint probability of  $Y_1$ ,  $Y_2$ ,  $R_1$  and  $R_2$  in  $M(G)$  is identifiable, then we say that the model  $G$  is  $YR$ -identifiable. In this paper, we only discuss decomposable graphical models. Figure (2) shows the eighteen possible decomposable graphical models.



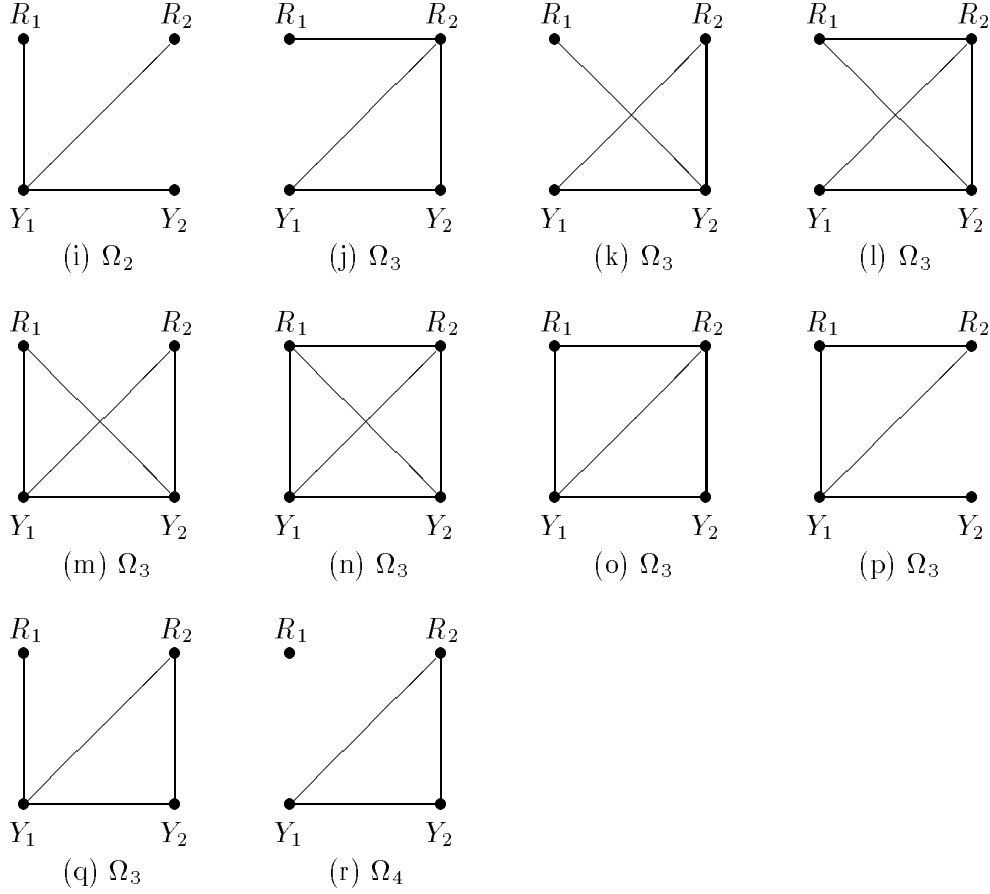


Figure 2. Eighteen composable graphs

These eighteen graphs are classified into four types  $\Omega_i$  for  $i = 1, \dots, 4$ . In  $\Omega_1$ , there are no edges between the response indicator  $R$  and the response variable  $Y$ . Thus the responses  $R_1$  and  $R_2$  does not depend on the true value of the response variables  $Y_1$  and  $Y_2$ . The mechanism of nonresponse (a) is *MCAR*, and the mechanism (b) is *MAR*. Both of them are ignorable. The parameter of these two models are identifiable.

**Theorem 1.** For the subgraphs in the set  $\Omega_2$ , we have that

1. any of graphical models (c), (d) and (h) is  $YR$ -identifiable if and only if  $Y_1 \not\perp\!\!\!\perp Y_2$ ;
2. all graphical models (e), (f) and (g) are  $YR$ -identifiable; and
3. the graphical model (i) is  $YR$ -identifiable if and only if  $Y_1 \not\perp\!\!\!\perp Y_2$  or  $R_2 \not\perp\!\!\!\perp Y_1$ .

There are eight observed data, but parameters in  $\Omega_3$  have more than eight degrees of

freedom. Thus they are nonidentifiable. In this situation, we can introduce some covariates to identify the parameters.

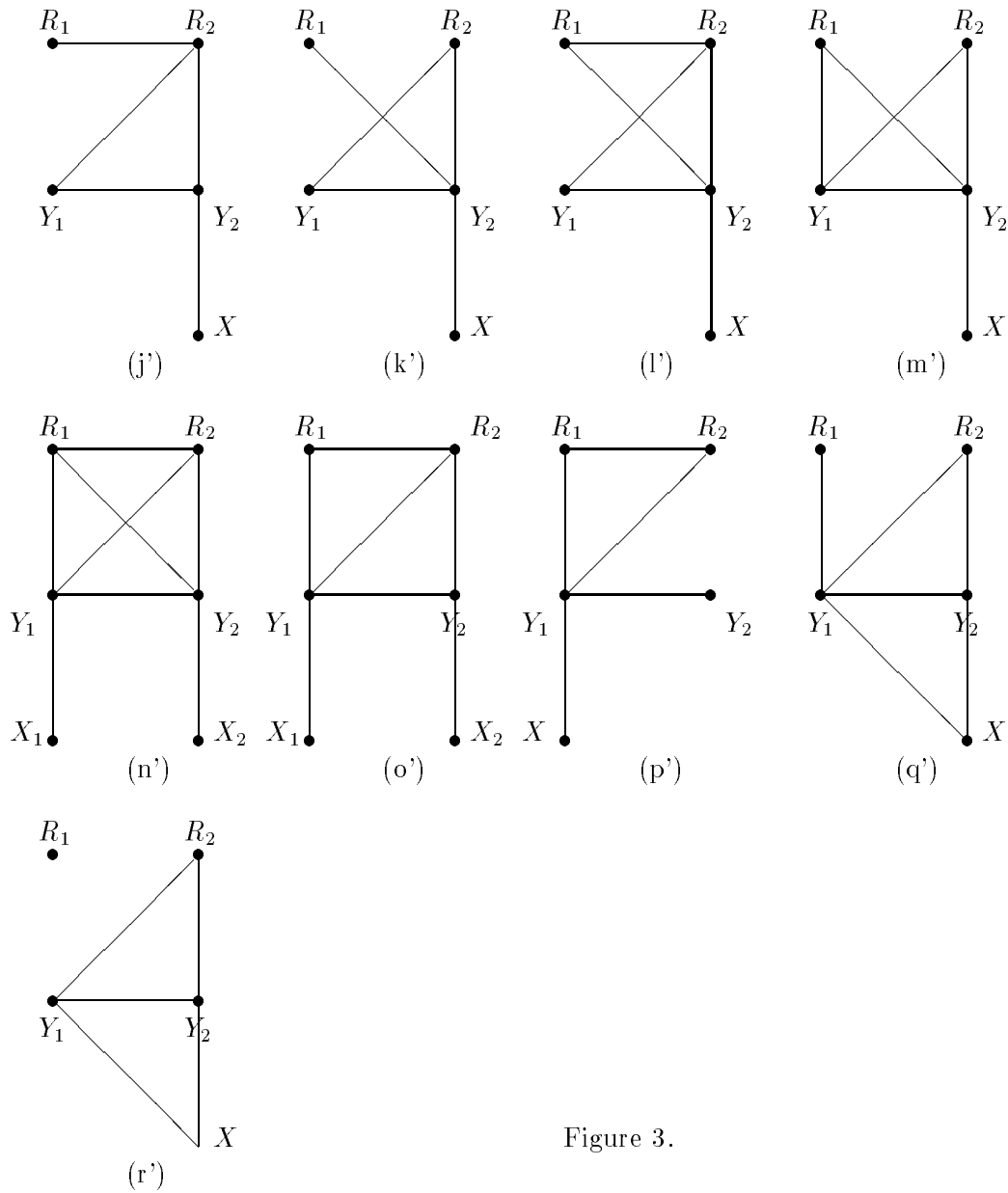


Figure 3.

For graphs (j), (k), (l) and (m), we introduce a covariate which is always observed with two values such that  $X \perp\!\!\!\perp (Y_1, R_1, R_2) | Y_2$ , as shown in Figure 3 (j'), (k'), (l') and (m').

**Theorem 2.** If  $X \not\perp\!\!\!\perp Y_2$ , then the graphical models (j'), (k'), (l') and (m') are  $YRX$ -identifiable (i.e.,  $P(y_1, r_1, y_2, r_2, x)$  is identifiable).

For graphs (n) and (o), we introduce a binary covariate  $X_1$  such that  $X_1 \perp\!\!\!\perp (Y_2, R_1, R_2) | Y_1$ , and another binary covariate  $X_2$  such that  $X_2 \perp\!\!\!\perp (Y_1, R_1, R_2) | Y_2$ , as shown in figure 3.

**Theorem 3.** If  $X_2 \not\perp\!\!\!\perp Y_2$  and  $X_1 \not\perp\!\!\!\perp Y_1$ , then the graphical models (n') and (o') are  $YRX_1X_2$ -identifiable (i.e.,  $P(y_1, r_1, y_2, r_2, x_1, x_2)$  is identifiable).

For graph (p), we introduce a always observed variable  $X$  with two values such that  $X \perp\!\!\!\perp (R_1, R_2, Y_2) | Y_1$ . We have graph (p') as shown in figure 3.

**Theorem 4.** If  $X \not\perp\!\!\!\perp Y_1$ , then the graphical model (p') is  $YRX$ -identifiable (i.e.,  $P(y_1, r_1, y_2, r_2, x)$  is identifiable).

For the graphs (q) and (r), we introduce a binary covariate  $X$  as shown in figure 3 (q') and (r'), where  $X$  satisfy the condition that  $X \perp\!\!\!\perp (R_1, R_2) | (Y_1, Y_2)$ .

**Theorem 5** If  $X \not\perp\!\!\!\perp Y_2 | Y_1$ , then the graphical models (q') and (r') are  $YRX$ -identifiable (i.e.,  $P(y_1, r_1, y_2, r_2, x)$  is identifiable).

An interest measure of association is the odds ratio in the margin table classified by  $Y_1$  and  $Y_2$  defined as

$$OR_{..} = \frac{P(Y_1 = 0, Y_2 = 0)P(Y_1 = 1, Y_2 = 1)}{P(Y_1 = 0, Y_2 = 1)P(Y_1 = 1, Y_2 = 0)}.$$

The odds ratio  $OR_{11} = P(Y_1 = 0, Y_2 = 0 | R_1 = 1, R_2 = 1)P(Y_1 = 1, Y_2 = 1 | R_1 = 1, R_2 = 1) / [P(Y_1 = 0, Y_2 = 1 | R_1 = 1, R_2 = 1)P(Y_1 = 1, Y_2 = 0 | R_1 = 1, R_2 = 1)]$  describes the association in the complete response subpopulation with  $R_1 = R_2 = 1$ . Assume that the complete response probability  $P(R_1 = R_2 = 1)$  is positive and larger than a given constant.

**Theorem 6** For the graphical models in  $\Omega_1, \Omega_2$  and the graph (p) in  $\Omega_3$ , we have that  $OR_{..} = OR_{11}$  and that  $\hat{OR}_{11} = n_{0011}n_{1111} / (n_{1011}n_{0111})$  is a strong consistent estimate of  $OR_{..}$ , where  $n_{ijkl}$  is the observed frequency of  $Y_1 = i, Y_2 = j, R_1 = k$  and  $R_2 = l$ .

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